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Title: The impact of obesity on the pharmacokinetics of drugs in adolescents and adults  
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Midazolam pharmacokinetics in morbidly obese patients following semi-simultaneous oral and intravenous administration: a comparison with healthy volunteers.

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ABSTRACT

Background
While \textit{in vitro} and animal studies have shown reduced CYP3A activity due to obesity, clinical studies in (morbidly) obese patients are scarce. As CYP3A activity may influence both clearance and oral bioavailability in a distinct manner, in this study the pharmacokinetics of the CYP3A substrate midazolam were evaluated after semi-simultaneous oral and intravenous administration in morbidly obese patients, and compared with healthy volunteers.

Methods
Twenty morbidly obese patients (mean body weight 144 kg (112-186 kg) and mean BMI 47 kg/m² (40-68 kg/m²) participated in the study. All patients received a midazolam 7.5 mg oral and 5 mg intravenous dose (separated by 159 ± 67 minutes) and per patient 22 samples over 11 hours were collected. Data from 12 healthy volunteers were available for a population pharmacokinetic analysis using NONMEM.

Results
In the three-compartment model in which oral absorption was characterized by a transit absorption model, population mean clearance (relative standard error %) was similar (0.36 (4.4%) L/min), while oral bioavailability was 60% (13.2%) in morbidly obese patients vs. 28% (7%) in healthy volunteers ($P < 0.001$). Central and peripheral volumes of distribution increased substantially with body weight (both $P < 0.001$) and absorption rate (transit rate constant) was lower in morbidly obese patients (0.057 (5%) vs. 0.130 (14%) min$^{-1}$, $P < 0.001$).

Conclusions
In morbidly obese patients, systemic clearance of midazolam is unchanged, while oral bioavailability is increased. Given the large increase in volumes of distribution, dose adaptations for intravenous midazolam should be considered. Further research should elucidate the exact physiological changes at intestinal and hepatic level contributing to these findings.
INTRODUCTION

The prevalence of obesity (body mass index (BMI) > 30 kg/m²) and morbid obesity (BMI > 40 kg/m²) is increasing rapidly. In 2010, 6.6% (15.5 million) of the U.S adult population was morbidly obese, a 70% increase since 2000 1, while 36% of the U.S. population was obese 2. In Europe, approximately 20% of the adult population is currently obese 3.

In obese mice, studies have shown reduced hepatic cytochrome P450 3A (CYP3A) protein expression 4-5. Similarly, in *in vitro* studies with hepatocytes from human fatty livers, reduced CYP3A expression and activity has been reported 6-7 with increasing severity of fatty liver and non-alcoholic steatohepatitis (NASH), which are both highly associated with (morbid) obesity 8-9. However, these measurements concern absolute values and were not normalized for the weight of the whole liver and/or the body weight of the mouse. CYP3A is an important enzyme system that is responsible for the primary metabolism of 25% of all clinically used drugs 10, including many drugs that are relevant for obese patients, such as statins, cardiovascular drugs, antipsychotics and oncolytic drugs 11. In obese as compared to non-obese subjects, it was shown that hepatic and intestinal CYP3A protein expression decreased with increasing BMI 12 and that oral clearance (CL/F) of CYP3A substrates such as triazolam, carbamazepine and taranabant was lower 13-14, even though a similar systemic clearance (CL) was found in obese individuals for midazolam 15. An explanation for the reduction in CYP3A activity upon obesity could be an increased state of inflammation, caused by infiltration of macrophages and adipocytes into the adipose tissue excreting inflammation markers and adipokines, such as interleukin (IL)-6 and tumor necrosis factor (TNF)-α 16-18. Both *in vitro* and animal studies have shown that inflammation factors such as IL-6 may decrease CYP3A expression, resulting in down regulation of CYP3A mediated metabolism 19-23. Finally, reduced CYP3A activity due to inflammation has also been shown in critically ill patients 24-25.

Midazolam which is considered a specific marker of CYP3A activity because it is primarily metabolized by CYP3A 26, is a widely applied oral or intravenous drug for sleeping disorders, (pre)anaesthesia, sedation for scopic interventions and in the intensive care unit. As CYP3A is located in both the liver and the intestines, the activity of the CYP3A enzyme is an important determinant of midazolam systemic clearance (CL) and oral bioavailability (F) 27-28. In view of the ever-increasing body weights of morbidly obese patients, in this pharmacokinetic study we evaluate the influence of morbid obesity on CYP3A-mediated systemic clearance and oral bioavailability of midazolam when studied after semi-simultaneous oral and intravenous administration, allowing these parameters to be characterized in a distinct manner. For the analysis, midazolam data from a healthy volunteer study with the same study design were also available. The results of this study are used to illustrate the consequences for dosing of oral and intravenous midazolam in morbidly obese patients.
METHODS

Study design and patients

This prospective observational study in morbidly obese patients (NTC01519726 and EudraCT 2011-003293-93) was approved by the local human research and ethics committee of the St. Antonius Ziekenhuis (VCMO, NL35861.100.11) and conducted in accordance with the principles of the Declaration of Helsinki and the Medical Research Involving Human Subjects Act (WMO) of the Netherlands. Before participation, all patients gave written informed consent.

Adult morbidly obese patients (BMI > 40 kg/m²) undergoing laparoscopic gastric bypass or sleeve surgery were eligible for inclusion in the study. Patients were excluded if they used CYP3A inducing or inhibiting medication according to Cytochrome P450 Drug Interaction Table 11, used products containing grapefruit, wild grape, banpeiyu, pomegranate, star fruit or black berry within 2 weeks before the study, were pregnant or breastfeeding, or suffered from renal insufficiency (estimated glomerular filtration rate (Modification of Diet in Renal Disease, MDRD4 < 60 mL/min)).

Study procedure

Twenty morbidly obese patients were studied on the day of laparoscopic bariatric surgery after an overnight fast. Midazolam was administered in a semi-simultaneous manner. Approximately 2.5 hours before induction of anaesthesia, patients received midazolam 7.5 mg orally as a tablet (Dormicum®, Roche). At the induction of anaesthesia (159 ± 67 minutes after the oral dose), 5 mg of intravenous midazolam (Midazolam, Actavis 5mg/mL) was administered. Blood samples were collected at T=0, 5, 15, 30, 45, 55, 65, 75, 90, 120, 150 minutes after the oral dose and T=5, 15, 30, 30, 90, 120, 150, 180, 210, 270, 330, 390, 510 minutes after the intravenous dose. Blood samples were collected in lithium heparin tubes and centrifuged at 3,000 rpm for 10 minutes at 4º C. Plasma was separated and immediately stored at -80 º C until analysis.

Blood samples to measure markers of liver function (aspartate aminotransferase (ASAT), alanine aminotransferase (ALAT), bilirubin, gamma-glutamyltransferase (gamma GT), and albumin) in morbidly obese patients were collected before the oral midazolam dose was administered.

Control group

Data from 12 healthy volunteers receiving midazolam in an identical semi-simultaneous dosing design were available for analysis (EudraCT 2009-010331-40). Subjects had to fast from 2 hours before drug administration and received a 2 mg oral midazolam solution (Midazolam, Synthon) at 10:00 am. After 150 minutes they received a 1 mg intravenous midazolam dose (Midazolam, Synthon) and blood samples were collected at T=15, 30,
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45, 60, 65, 70, 75, 80, 90, 120, 148, 155, 165, 180, 210, 240, 270, 330, 390 minutes after oral dose.

Drug assay

In the plasma samples from morbidly obese patients, midazolam was analyzed using a MassTox® TDM series A BASIC-Kit for ultra performance liquid chromatography (UPLC)-tandem mass spectrometry (MS/MS) from Chromsystems Instruments & Chemicals GmbH (Gräfelfing/München, Germany), a commercially available kit including mobile phases, dilution buffers and extraction buffer. For sample preparation, 25 μL of MassTox® BASIC-Kit A extraction buffer was added to 50 μL of the sample. After short vortex mixing, this mixture was let to incubate for 2 minutes. Then 250 μL MassTox® BASIC-Kit A internal standard mix was added and the mixture was vortexed for 30 seconds and centrifuged at 15,000 g for 5 minutes. Ten μL of the supernatant was injected onto the ULPC column (Chromsystems MasterColumn) without a precolumn using a Waters Acquity UPLC system connected to a Waters TQD mass spectrometer with electrospray ionization. The column was kept at 25º C. Eluent was used at a flow rate of 0.6 mL/min. Intra-assay and inter-assay coefficients of variation were 4.7% and 3.3%, respectively. Midazolam recovery was 90%. The lower limit of quantification (LLOQ) was 0.8 ng/mL.

In the plasma samples from healthy volunteers, midazolam was measured using a validated liquid chromatographic-tandem mass spectrometric (LC-MS/MS) assay. Briefly, 500 μL acetonitrile containing midazolam-D4 (4 μg/L) was added to 200 μL serum. After 3 min vortex mixing and 5 min centrifugation at ambient temperature the supernatant was collected and transferred into an autosampler vial. Next, 10 μL was injected on an Atlantis T3 C18 3µm column (2.1 x 50 mm; Waters), protected with a guard column (ODS; 4 x 3 mm), which was kept at 30 ºC. Gradient elution was performed with a mobile phase consisting of 0.1% v/v aqueous formic acid and 0.1% v/v formic acid in acetonitrile at a flow rate of 0.3 mL/min. The effluent was monitored with a Micromass Quattro Micro triple-quadrupole mass spectrometric detector (Waters). The detector was operated in the positive electrospray ionization mode and configured in the multiple reaction monitoring (MRM) mode. Within-day and between-day inaccuracy and imprecision were less than 5%. The LLOQ was 0.3 ng/mL.

Population pharmacokinetic analysis and internal validation

Population pharmacokinetic modelling was performed on all data by means of nonlinear mixed effects modelling using NONMEM® (version 7.2; GloboMax LLC, Hanover, MD, USA). Pirana (2.7.1; Pirana Software & Consulting BV) and R (2.15) were used to visualize the data. All midazolam plasma concentration values that were received from the laboratory were inserted in the datafile, even if these were below the limit of quantification (LOQ).
Of the 434 samples from morbidly obese patients, 42 were below the LOQ. For healthy volunteers, no data was below the LOQ.

Discrimination between different models was made by comparison of the objective function value (OFV, i.e. -2 log likelihood (-2LL)). A P value below 0.05, representing a decrease of 3.84 in the OFV for one degree of freedom, was considered statistically significant. In addition, goodness-of-fit plots (observed vs. individual-predicted concentrations, observed vs. population-predicted concentrations, conditional weighted residuals vs. time and conditional weighted residuals vs. population-predicted concentrations plots) were used for diagnostic purposes. Furthermore, the confidence interval of the parameter estimates, the correlation matrix and visual improvement of the individual plots were used to evaluate the model. The internal validity of the population pharmacokinetic model was assessed by the bootstrap re-sampling method using 500 replicates and normalized prediction distribution errors (NPDE) using 1,000 simulation of each dataset. Parameters obtained with the bootstrap replicates were compared with the estimates obtained from the original dataset. NPDE plots were checked for normal distribution characteristics and trends in the data errors.

For the structural model, one-, two, and three-compartment models with an oral dosing compartment were tested. To describe the midazolam oral absorption phase, zero order and first order absorption models were tested, in addition to a lag time model and a transit absorption model. For the transit absorption model, a varying number of transit compartments was tested. As the transit rate ($K_t$) was set equal to the absorption rate ($K_a$), the mean transit time (MTT) can be calculated from $K_t$ with $(n+1)/K_t$ in which $n$ is the number of transit compartments.

For the statistical model, the individual parameter estimate (empirical Bayes estimate or post hoc value) of the $i^{th}$ individual was modelled according to (Equation 1):

$$\theta_i = \theta_{\text{mean}} \times e^{\eta_i}$$

(Eq. 1)

where $\theta_{\text{mean}}$ is the population mean, and $\eta_i$ is a random variable for the $i^{th}$ individual with a mean of zero and variance of $\omega^2$, assuming log-normal distribution in the population.

For residual variability, resulting from assay errors, model misspecifications and other unexplained sources, a proportional error model and a combined proportional and additive model was tested for each of the data sets. The $j^{th}$ observed midazolam concentration of the $i^{th}$ healthy volunteer ($Y_{ij}$) is described by Equation 2, while the $j^{th}$ observed midazolam concentration of the $i^{th}$ morbidly obese patient ($Y_{ij}$) is described by Equation 3:

$$Y_{ij} = C_{\text{pred},ij} \times (1+\epsilon_{ij})$$

(Eq. 2)
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\[ Y_{ij} = C_{\text{pred},ij} \times (1 + \varepsilon_{ij}) + \varepsilon_{ij} \]  

(Eq. 3)

where \( C_{\text{pred},ij} \) is the population predicted midazolam concentration of the \( i^{\text{th}} \) individual at the \( j^{\text{th}} \) time, and \( \varepsilon_{ij} \) is a random variable with a mean of zero and variance of \( \sigma^2 \).

**Covariate analysis**

Covariates were plotted independently against the individual post hoc values and eta estimates of pharmacokinetic parameters to visualize potential relations. The following covariates were tested: total body weight (TBW), BMI, lean body weight (LBW) \(^{33}\), sex, morbid obesity and age. All covariates except for sex and morbid obesity were tested using linear and allometric equations (Equation 4 and 5):

\[ P_i = P_p \times (1 + W \times (\text{COV}_i - \text{COV}_{\text{median}})) \]  

(Eq. 4)

\[ P_i = P_p \times \left( \frac{\text{COV}_i}{\text{COV}_{\text{median}}} \right)^X \]  

(Eq. 5)

where \( P_i \) and \( P_p \) represent individual and population parameter estimates, respectively; \( \text{COV} \) represents the covariate; \( \text{COV}_{\text{median}} \) represents the median value of the covariate for the population; \( W \) represents a correlation factor between the population pharmacokinetic parameters and the change in covariate value; and \( X \) represents the exponential scaling factor for a power function. The binary covariates ‘sex’ and ‘morbid obesity’ were tested using Equation 6:

\[ P_i = P_p \times Z^{\text{COV}} \]  

(Eq. 6)

where \( P_i \) and \( P_p \) represent individual and population parameter estimate, \( Z \) the estimated factor of increase or decrease for the patients subgroup with \( \text{COV} \) equalling one.

Potential covariates were separately entered into the model and statistically tested by use of the OFV and, if applicable, the 95% confidence interval of the additional parameter. In addition, if applicable, it was evaluated whether the inter-individual variability (eta) in the parameter concerned decreased upon inclusion of the covariate on the parameter and whether the trend in the eta vs. covariate plot had resolved. When more than one significant covariate for the simple model was found, the covariate-adjusted model with the largest decrease in the OFV was chosen as a basis to sequentially explore the influence of additional covariates with the use of the same criteria. Finally, after forward inclusion \( (P < 0.05) \), a backward exclusion procedure was applied to justify the inclusion of a covariate \( (P < 0.01) \). The choice of the covariate model was further evaluated as discussed in the Population pharmacokinetic analysis and internal validation section.
Model simulations
Using NONMEM® 7.2, the final population pharmacokinetic model was used to simulate the concentration-time profiles of four typical patients from the dataset, including one healthy volunteer of 76 kg and three morbidly obese patients of 112, 145 and 186 kg. The 76 kg and 145 kg dose simulations represent the median body weight of the healthy volunteer and morbidly obese patient group, respectively. In addition, the 112 kg and 186 kg dose simulations represent the extremes of the body weight range of the morbidly obese patient group (see Table 1).

RESULTS

Patients and data
Twenty morbidly obese patients participated in this study and a mean of 22 ± 3 samples per patient were available for analysis. In addition, data from 12 healthy volunteers with 19 midazolam concentrations per subject were used as a control group in this analysis.

Liver function markers in morbidly obese subjects were all within three times the upper limit of normal, with the vast majority being within two times the upper limit of normal of the different markers. The demographics of all subjects are summarized in Table 1.

Table 1 Characteristics of 20 morbidly obese patients and 12 healthy volunteers.

<table>
<thead>
<tr>
<th></th>
<th>Morbidly obese patients (n=20)</th>
<th>Range</th>
<th>Healthy volunteers (n=12)</th>
<th>Range</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Female/Male</td>
<td>12/8</td>
<td>0/12</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age (years)</td>
<td>43.6 ± 7.6</td>
<td>26 - 57</td>
<td>22.0 ± 3.1</td>
<td>18 - 27</td>
<td>P &lt; 0.001</td>
</tr>
<tr>
<td>Body weight (kg)</td>
<td>144.4 ± 21.7</td>
<td>112 - 186</td>
<td>76.0 ± 8.7</td>
<td>63 - 93</td>
<td>P &lt; 0.001</td>
</tr>
<tr>
<td>Lean body weight (kg)</td>
<td>71.5 ± 11.9</td>
<td>53 - 95</td>
<td>61.2 ± 5.0</td>
<td>53 - 70</td>
<td>P &lt; 0.005</td>
</tr>
<tr>
<td>BMI (kg/m²)</td>
<td>47.1 ± 6.5</td>
<td>40 - 68</td>
<td>22.3 ± 2.4</td>
<td>19 - 26</td>
<td>P &lt; 0.001</td>
</tr>
<tr>
<td>Number of samples per patient</td>
<td>21.7 ± 2.7</td>
<td>13 - 24</td>
<td>19 ± 0.0</td>
<td>19 - 19</td>
<td>P &lt; 0.001</td>
</tr>
<tr>
<td>Samples below the limit of quantification (%)</td>
<td>9.7</td>
<td>-</td>
<td>0</td>
<td>-</td>
<td>-</td>
</tr>
</tbody>
</table>

Values are expressed as mean ± standard deviation (range) unless specified otherwise

Population pharmacokinetic model and validation
A three-compartment model with two equalized peripheral volumes of distribution best fitted the data. This model showed an improved fit over a two-compartment model, while a full three compartment model could not be estimated with adequate precision. The pharmacokinetic model was parameterized in terms of the oral Ka, oral bioavailability (F), volume of distribution of the central compartment (V1), two equalized volumes of distribu-
tion of the peripheral compartments (V2 and V3), inter-compartmental clearances from the central compartment to each peripheral compartment (Q and Q2), and clearance from the central compartment (CL). Midazolam oral absorption described by three transit absorption compartments proved superior over the other oral absorption models (zero and/or first order absorption models or a lag time model) (Figure 1). An omega block was implemented to account for the correlation between the central and peripheral volume of distribution. Table 2 shows the parameter estimates of the simple model without covariates.

**Figure 1** Schematic illustration of the population pharmacokinetic model of oral and intravenous administered midazolam. CL= clearance, F= oral bioavailability, K_a= oral absorption rate, K_{tr}= transit rate constant, MDZ= midazolam, Q= intercompartmental clearance to first peripheral volume, Q2= intercompartmental clearance to second peripheral volume, V= volume of distribution.

In the covariate analysis, a significant influence of TBW or ‘morbid obesity’ was found on four different parameters, which is visualized in Figure 2 where the post hoc parameters of the simple model without covariates are given. It was found that the peripheral volumes of distribution increased in a non-linear manner with TBW (\(P < 0.001\), -24 ΔOFV), and that the central volume of distribution showed a linear increase with body weight (\(P < 0.001\), -17 ΔOFV). For oral bioavailability and absorption rate (or K_{tr}), ‘morbid obesity’
was a significant covariate and significantly improved the model \( P < 0.001, -22 \Delta OFV \) and \( P < 0.001, -20 \Delta OFV \), respectively. For clearance, there was a trend towards a positive influence of LBW but not for TBW; however, the statistical significance was insufficient for inclusion of LBW in the final covariate model \( P < 0.05, -4 \Delta OFV \), the estimated correlation factor was not estimated with adequate precision and the eta for clearance value was not reduced. Eta distributions for clearance vs. body weight and LBW of the simple and final covariate model are included in the Supplementary Material (Supplementary Figure 1, see end of this Chapter). After inclusion of the covariates in the model, the trends in eta value of the parameter and the covariate had disappeared and no residual trends were observed (Supplementary Figure 2). All parameter estimates of the final model are shown in Table 2. Figure 3 demonstrates the goodness-of-fit plots of the final covariate model.

![Figure 2](image_url)

**Figure 2** Post hoc parameter estimates of morbidly obese individuals \((n=20, \text{black dots})\) and healthy individuals \((n=12, \text{grey dots})\) from the simple pharmacokinetic model vs. total body weight, including central volume of distribution vs. total body weight \((a)\), peripheral volume of distribution vs. total body weight \((b)\), oral bioavailability vs. total body weight \((c)\), and oral absorption rate vs. total body weight \((d)\).
### Table 2: Population pharmacokinetic parameters of the simple and final pharmacokinetic model for midazolam in 20 morbidly obese patients and 12 healthy volunteers and results from a bootstrap analysis of the final model (479/500 resamples successful).

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Simple model (RSE%)</th>
<th>Final model (RSE%)</th>
<th>Bootstrap (SE)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Fixed effects</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>CL (L/min)</td>
<td>0.36 (4.9)</td>
<td>0.359 (4.4)</td>
<td>0.358 (0.016)</td>
</tr>
<tr>
<td>F</td>
<td>0.414 (12.8)</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>F morbidly obese</td>
<td>-</td>
<td>0.603 (13.2)</td>
<td>0.603 (0.081)</td>
</tr>
<tr>
<td>F healthy volunteers</td>
<td>-</td>
<td>0.284 (7.0)</td>
<td>0.286 (0.020)</td>
</tr>
<tr>
<td>Ka (min⁻¹) = Ktr</td>
<td>0.086 (3.4)</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>Kₐ = Kᵣ, morbidly obese</td>
<td>-</td>
<td>0.057 (13.6)</td>
<td>0.058 (0.059)</td>
</tr>
<tr>
<td>Kₐ = Kᵣ, healthy volunteers</td>
<td>-</td>
<td>0.13 (5.1)</td>
<td>0.130 (0.006)</td>
</tr>
<tr>
<td>V.central (L)</td>
<td>36.4 (7.8)</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>V.central, 127 kg</td>
<td>-</td>
<td>44.1 (16.1)</td>
<td>43.6 (6.9)</td>
</tr>
<tr>
<td>Z</td>
<td>-</td>
<td>0.0105 (15.8)</td>
<td>0.0102 (0.002)</td>
</tr>
<tr>
<td>V.midazolam peripheral (L)</td>
<td>76.6 (8.4)</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>V.peripheral, 127 kg</td>
<td>-</td>
<td>139 (15.2)</td>
<td>138.7 (22.9)</td>
</tr>
<tr>
<td>W</td>
<td>-</td>
<td>3.06 (8.2)</td>
<td>3.07 (0.28)</td>
</tr>
<tr>
<td>Q (L/min)</td>
<td>1.31 (12.8)</td>
<td>1.33 (11.8)</td>
<td>1.33 (0.143)</td>
</tr>
<tr>
<td>Q2 (L/min)</td>
<td>0.153 (12.1)</td>
<td>0.15 (14.6)</td>
<td>0.15 (0.023)</td>
</tr>
<tr>
<td><strong>Interindividual variability (%)</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>CL</td>
<td>19.7% (32.6)</td>
<td>18.1% (30.7)</td>
<td>17.2% (14.6)</td>
</tr>
<tr>
<td>F</td>
<td>61.2% (20.8)</td>
<td>26.4% (17.4)</td>
<td>25.4% (14.6)</td>
</tr>
<tr>
<td>Kₐ = Kᵣ</td>
<td>50.7% (10.9)</td>
<td>41.4% (12.8)</td>
<td>39.9% (18.4)</td>
</tr>
<tr>
<td>V.central</td>
<td>102.8% (13.2)</td>
<td>55.2% (17.5)</td>
<td>53.5% (30.8)</td>
</tr>
<tr>
<td>V.peripheral</td>
<td>152.3% (13.9)</td>
<td>34.4% (26.5)</td>
<td>33.6% (23.7)</td>
</tr>
<tr>
<td>Correlation between eta V.central and V.peripheral</td>
<td>0.783 (50.0)</td>
<td>0.12 (24.5)</td>
<td>0.10 (0.058)</td>
</tr>
</tbody>
</table>

### Residual variability (%)

<p>| | | | |</p>
<table>
<thead>
<tr>
<th></th>
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</tr>
</thead>
<tbody>
<tr>
<td>Proportional error healthy volunteers</td>
<td>10.0% (21.5)</td>
<td>10.0% (21.3)</td>
<td>9.9% (4.4)</td>
</tr>
<tr>
<td>Proportional error morbidly obese patients</td>
<td>31.0% (17.1)</td>
<td>46.7% (11.6)</td>
<td>46.6% (15.6)</td>
</tr>
<tr>
<td>Additive error morbidly obese patients</td>
<td>3.1 (37.3)</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>OFV</td>
<td>4077</td>
<td>4003</td>
<td>3982 (145)</td>
</tr>
</tbody>
</table>

CL = systemic clearance of midazolam, F = oral bio-availability, Kᵣ = oral absorption rate, Kᵣ = transit compartment rate, Vᵣ = volume of distribution, Qᵣ = inter-compartmental clearance of midazolam between central and first peripheral compartment, Q₂ᵣ = inter-compartmental clearance of midazolam between central and second peripheral compartment, OFV = Objective Function Value, RSE = relative standard error, SE = standard error, TBW = total body weight, Vᵣ = volume of distribution

* Mean transit time is 46.5 minutes.

* Mean transit times are 70.2 and 30.8 minutes, respectively.
The plots show no remaining bias for predicted midazolam concentrations in morbidly obese patients or healthy volunteers. The final model was internally validated by means of 500 bootstrap runs (Table 2), which were successful in 96% of the runs and confirmed the parameter values. Finally, an NPDE analysis was performed showing a normal distribution of errors without trends, except for the initial midazolam concentrations in morbidly obese patients, which were below the LOQ (Supplementary Figure 3). The NONMEM® control file of the final model is provided as Supplementary Material (Supplementary File 1).

**Figure 3** Observed vs. individual predicted midazolam concentrations (a), observed vs. population predicted midazolam concentrations (b), conditional weighted residuals (CWRES) vs. population predicted midazolam concentrations (c) and conditional weighted residuals (CWRES) vs. time (d) of the final model for 20 morbidly obese patients (black dots) and 12 healthy volunteers (grey dots). The dashed line represents the line of identity (x=y).
Simulations

Figure 4 shows population predicted midazolam concentrations after a 5 mg intravenous bolus dose, a 2.5 mg/h continuous infusion for a duration of 10 days (14,400 minutes) and a 7.5 mg oral dose in 4 typical subjects from the dataset (i.e. 76 kg, 112, 145 and 186 kg). The plot shows that for the intravenous bolus dose, midazolam (peak) concentrations ($C_{max}$) are lower in morbidly obese patients (Figure 4a), which may be the result of the higher central volume of distribution in morbidly obese patients. In addition, the plot illustrates the longer midazolam half-life ($t_{1/2}$) in morbidly obese patients, which can be attributed to the increase in volumes of distribution with body weight, while clearance is the same in all patients. The continuous intravenous infusion simulation (Figure 4b) shows that with increasing body weight the midazolam steady-state concentrations are reached at a later time point. In a 76 kg healthy volunteer, steady-state is expected to be reached after 24 hours, while this is 170 hours in a 145 kg morbidly obese patient and >240 hours in a 186 kg morbidly obese patient (Figure 4b). Finally, the oral mid-
azolam dose simulations (Figure 4c) show that the time to reach the $C_{\text{max}}$ is later (31 minutes) in morbidly obese patients than in healthy volunteers, while the $C_{\text{max}}$ is slightly lower in morbidly obese patients.

**DISCUSSION**

As there is only limited information on the influence of morbid obesity on CYP3A-mediated clearance of drugs in patients, this study aimed to evaluate the pharmacokinetics of CYP3A substrate midazolam in morbidly obese patients following oral and intravenous midazolam administration. As clearance after oral dosing is dependent on oral bioavailability, which may be influenced by CYP3A enzyme activity in the intestines, this semi-simultaneous design allows for an estimation of both systemic clearance and oral bioavailability in a distinct manner. An available dataset of midazolam concentrations collected on the basis of an equivalent study design in healthy volunteers allowed for a head-to-head comparison between morbidly obese patients and non-obese healthy subjects. The results from this study show that midazolam clearance was similar in morbidly obese patients and healthy volunteers, oral bioavailability was substantially higher (60% instead of 28%), oral absorption rate was reduced and that the central and peripheral volumes of distribution increased substantially with body weight. Particularly for intravenous dosing, the net results of all these changes should be considered when administering midazolam to morbidly obese patients.

In this study, we could not identify an influence of morbid obesity on the systemic clearance (CL) of midazolam, even though a wide range in body weights was included in this study. We did find a trend of increasing midazolam clearance with lean body weight; however, this trend was not strong enough for inclusion in the final covariate model. Possibly, the patient numbers in this analysis (n=12 + n=20) are insufficient to adequately detect a small increase in clearance with LBW. While these results indicate a lack of change in absolute hepatic CYP3A mediated metabolism of midazolam in morbidly obese individuals, the results are in contrast with our expectations of a lower midazolam clearance in morbidly obese patients which was based on reports in *in vitro* and animal studies 4-7 and on oral clearance of CYP3A substrates in studies in obese subjects 13-14. Assuming that indeed the relative CYP3A activity per unit of liver is reduced in morbidly obese patients 12, we hypothesize that this effect may be counteracted by a higher liver volume 34, resulting in a similar absolute hepatic CYP3A metabolising capacity in both groups. In agreement with this hypothesis, also Greenblatt et al. found no significant difference in absolute systemic clearance of midazolam between normal weight (66 ± 2 kg) and obese subjects (117 ± 8 kg) (0.53 ± 0.04 vs. 0.47 ± 0.04 L/min, respectively) 15. However for a study with triazolam, another benzodiazepine CYP3A substrate, a lower
oral clearance (CL/F) was found for obese patients. Lower CL/F values were also found for obese patients for the CYP3A substrates taranabant and carbamazepine. Based on the results found in the current midazolam study upon both oral and intravenous administration, it may be hypothesized that these lower oral clearance (CL/F) values are due to an increase in oral bioavailability (F) instead of a decrease in systemic clearance (CL).

Oral bioavailability was found to be higher in morbidly obese patients compared to healthy volunteers (0.60 (13.2%) vs. 0.28 (7.0%). In contrast, Greenblatt et al. found similar values of oral midazolam bioavailability in obese (0.42 ± 0.04) and normal weight patients (0.40 ± 0.03) (P > 0.05). This disagreement in results may be explained by the higher body weights of the morbidly obese subjects in our study vs. the study of Greenblatt et al. (mean of 144 ± 22 vs. 117 ± 8 kg). In addition, the concentration-time profiles after oral and intravenous midazolam of a non-obese and an obese subject shown in their publication, may also point at a higher bioavailability in the obese patient. We anticipate that the increase in oral bioavailability in morbidly obese patients found in our study may be due to reduced CYP3A metabolizing activity in the intestines. Ulvestad et al. found in a study with 19 obese individuals (median BMI 45 (34-59) kg/m²) that CYP3A4 protein expression in the small intestine and liver is lower with increasing BMI. Another possible cause of increased bioavailability is an increase in splanchnic blood flow, which has been reported before in morbidly obese patients. An increase in villous blood flow in the gut wall will cause an increase in substrate transport and thus carry the substrate away from the intestinal CYP3A metabolizing enzymes. Moreover, increased intestinal permeability may be responsible for increased midazolam bioavailability as obese patients showed increased paracellular absorption measured with lactulose and chromium (Cr)-EDTA, which may possibly be due to reduced tight junction function. A question would be whether the observed difference in absorption rate (Ka) between the morbidly obese patients and healthy volunteers, which can be attributed to a difference in formulation (tablet vs. oral solution), may have contributed to the reported difference in oral bioavailability. In our opinion, this difference in formulation is unlikely to influence oral bioavailability, as midazolam is a highly soluble and permeable drug which is expected to be 100% absorbed in the intestines. Correspondingly, a study in which 6 healthy volunteers received a 10 mg midazolam oral solution, a 10 mg tablet and a 5 mg intravenous bolus dose, showed similar oral bioavailability after both oral dose formulations, 0.35 ± 0.07 vs. 0.38 ± 0.12 (P > 0.6), indicating no influence of oral formulation on midazolam bioavailability.

To understand the net result of the influence of different degrees of (morbid) obesity on each of the midazolam pharmacokinetic parameters, simulations using the final model were performed to yield midazolam concentration-time profiles for subjects with different body weights. The dose simulations show that the same intravenous
bolus dose to all subjects leads to lower initial concentrations in morbidly obese patients due to a substantially higher central and peripheral volume of distribution. This observed increase in volume of distribution is in agreement with the midazolam study of Greenblatt et al. in which a substantial increase in total volume of distribution for obese vs. normal weight subjects of $311 \pm 27$ L vs. $114 \pm 7$ L ($P < 0.001$) was also reported. Potentially, these results can be explained by an increase in body volume in terms of both well-perfused compartments (organ and blood volume) and adipose tissue with obesity, which is of specific relevance because midazolam is a lipophilic drug. As such, directly after an intravenous bolus dose, lower midazolam concentrations and associated effects may be expected in morbidly obese patients. In addition, morbidly obese patients show an increased elimination half-life (Figure 4), which can be attributed to the larger volumes of distribution as well, as they allow for significant midazolam disposition from the blood and may lead to prolonged midazolam effects in morbidly obese patients vs. healthy volunteers. In contrast, a similar midazolam oral dose will result in only slightly lower initial concentrations in morbidly obese patients vs. healthy volunteers because the increased oral bioavailability counteracts the influence of increased central volume of distribution on midazolam peak concentrations (Figure 4). Finally, the increase in the volumes of distribution with body weight also explains the increased duration for morbidly obese patients to reach steady-state concentrations after a continuous intravenous infusion. This phenomenon has been described before for diazepam in obese patients. Therefore a loading dose or a higher initial continuous infusion rate may be considered to reach midazolam steady-state concentrations more rapidly in morbidly obese patients.

There are some limitations to this study. Firstly, the sampling duration after oral administration may have been relatively short. Particularly in morbidly obese patients, the midazolam peak concentrations after the oral dose occurred at approximately 90 minutes post dose, leaving only a 60 minute time interval to collect data on the concentration decline after oral dose before the intravenous dose was administered. However, in six of the 20 patients this interval was > 180 minutes (due to a delay in the surgery schedule), thus providing significant information on the midazolam pharmacokinetics after oral absorption in the morbidly obese patients. Secondly, the healthy volunteer group lacks a late sample post intravenous dose, which have may have an effect on the clearance and peripheral volume of distribution estimates of the healthy volunteers and thus obscure the covariate analysis. However, estimated pharmacokinetic parameter values for this group closely match those found in previous midazolam pharmacokinetic studies in healthy volunteers, indicating adequate precision of the pharmacokinetic parameters in healthy volunteers and justifying the results from the covariate analysis. Thirdly, morbidly obese patients underwent surgery during the study, which may influence midazolam clearance and distribution. However, we think that surgery was only
of minor influence as only the intravenous dose was administered during surgery and systemic clearance and volume of distribution found in this study were fairly similar to earlier reported values in non-surgery obese patients. Finally, the stable isotope method for determining oral bioavailability in a single person on a single occasion may have been preferable over the current semi-simultaneous dosing design. Though, the semi-simultaneous oral-intravenous administration method has proved a reliable and accurate method for estimating oral bioavailability and systemic clearance in a single person, on a single occasion as well. Moreover, the available control data (midazolam concentrations in healthy volunteers) was gathered in a semi-simultaneous design. Lastly, the preparation of the labelled drug and the determination of the labelled drug in the samples is very expensive and labour intensive. For these reasons, we have chosen to apply the semi-simultaneous design.

In conclusion, this study shows that midazolam hepatic clearance was not changed in morbidly obese patients vs. healthy volunteers, while oral bioavailability was increased in morbidly obese patients. Midazolam central and peripheral volumes of distribution increased substantially with body weight resulting in lower midazolam concentrations after intravenous bolus administration and in increased duration to reach steady state concentrations after midazolam continuous infusion in morbidly obese patients in comparison to healthy volunteers. Finally, initial midazolam concentrations after an oral dose were similar in morbidly obese patients vs. healthy volunteers. Further research should elucidate the exact physiological changes at intestinal and hepatic level contributing to these findings.

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3. IASO. International Association for the Study of Obesity. In: IASO.


**Supplementary Material**

**Supplementary Figure 1** Eta distributions for clearance for the simple (a and c) and final (b and d) pharmacokinetic model vs. lean body weight (a and b) and body weight (c and d) in 20 morbidly obese patients (black dots) and 12 healthy volunteers (grey dots).
Supplementary Figure 2 Eta distributions for central volume of distribution (a and b), peripheral volume of distribution (c and d), oral bioavailability (e and f), oral absorption rate (Ktr) (g and h) in 20 morbidly obese patients (black dots) and 12 healthy volunteers (grey dots) for the simple (a, c, e, g) and final midazolam pharmacokinetic model (b, d, f, h).
Supplementary Figure 3. Normalized prediction distribution errors (npde) for the midazolam data from 20 morbidly obese patients (a) and from 12 healthy volunteers (b). The X indicates the time after dose and Y indicates the midazolam concentration in nmol/L.
Supplementary File 1  NONMEM control stream of the final pharmacokinetic model.

$PROBLEM Final model #61
$INPUT ID TIME AMT RATE DROP DV CMT MDV TAD TADI OBES TBW LBW IBW BMI WHR AGE SEX LOQ IV ROUN
; Parent=run18
$DATA BOTH.occ1.incl.BLOQ.11.csv
$SUBROUTINES ADVAN6 TOL=5
$MODEL
COMP=(PODOSE)
COMP=(CENTRAL)
COMP=(PERIP)
COMP=(PERIP2)
COMP=(TRANSIT1)
COMP=(TRANSIT2)
COMP=(TRANSIT3)
$PK
TVCL= THETA(1)
CL= TVCL*EXP(ETA(1))

TVF1= THETA(2)*THETA(10)**OBES
F1= TVF1*EXP(ETA(2)) ; oral bioavailability

TVV2= THETA(3)*(1+THETA(8)*(TBW-127))
V2= TVV2*EXP(ETA(5)) ; volume (L)

TVQ= THETA(4)
Q= TVQ*EXP(ETA(3)) ; intercompartmental clearance (L/min)

TVV3= THETA(5)*(TBW/127)**THETA(9)
V3= TVV3*EXP(ETA(6)) ; peripheral volume of distribution (L)

TVKA= THETA(6)*THETA(11)**OBES
KA= TVKA*EXP(ETA(4)) ; oral absorption rate or Ktr (min-1)

KTR= KA
V4= V3 ; 2nd volume periph. compartment (L)
Q4= THETA(7) ; clearance to 2nd periph. compartment (L/min)
;
S2=V2
S3=V3
;
K15=KA
K56=KTR
K67=KTR
K72=KTR
K20=CL/V2
K23=Q/V2
K32=Q/V3
K24=Q4/V2
K42=Q4/V4

$DES
DADT(1)= -K15*A(1)
DADT(2)= KTR*A(7) -K23*A(2) +K32*A(3) -K24*A(2) +K42*A(4) -K20*A(2)
DADT(3)= K23*A(2) -K32*A(3)
DADT(4)= K24*A(2) -K42*A(4)
DADT(5)= K15*A(1) -KTR*A(5)
DADT(6)= KTR*A(5) -KTR*A(6)
DADT(7)= KTR*A(6) -KTR*A(7)

$ERROR
COM1=0
IF (OBES.EQ.0) COM1=1
COM2=0
IF (OBES.EQ.1) COM2=1

IPRED=F ; individual prediction

Y1=IPRED*(1+ERR(1)) ; healthy volunteers
Y2=IPRED*(1+ERR(2)) ; morbidly obese patients

Y=Y1*COM1+Y2*COM2

IRES=DV-IPRED ; individual residual
DEL=0 ; if F=0, than see next lines
IF(IPRED.EQ.0) DEL=1
IWRES=(1-DEL)*IRES/(IPRED+DEL)

$THETA
(0, 0.2) ; CL (L/min)
(0, 0.3) ; F1 healthy volunteers
(0, 30) ; V2 (L)
(0, 0.9) ; Q (L/min)
(0, 100) ; V3 (L)
(0, 0.1) ; KA healthy volunteers
(0, 0.1) ; Q4
(0, 0.003) ; TBW V2
(0, 2) ; TBW V3
(0, 2) ; fraction F1 morbidly obese
(0, 0.5) ; fraction KA morbidly obese

$OMEGA
0.03 ; CL
0.1 ; F1
0 FIX ; Q
0.2 ; KA

$OMEGA BLOCK(2)
0.3 ; V2 (central)
0.05 0.1 ; V3 (peripheral)

; $SIGMA
0.005 ; healthy volunteers
0.1 ; morbidly obese

$EST SIGDIG=3 MAXEVAL=9999 PRINT=5 NOABORT METHOD=1 INTERACTION POSTHOC
$COV COMP
$TABLE ID TIME TAD TADI IPRED IWRES CWRES CMT ROUN OBES IV NOPRINT ONEHEADER
FILE=sdtab61
$TABLE ID TAD CMT TADI ROUN OBES IV CL V2 Q V3 Q4 KA F1 V4 TBW LBW IBW AGE
BMI WHR ETA(1) ETA(2) ETA(3) ETA(4) ETA(5) ETA(6) NOPRINT NOAPPEND ONEHEADER
FILE=patab61
$TABLE ID TAD TADI TBW LBW IBW BMI WHR AGE SEX NOPRINT NOAPPEND ONEHEADER
FILE=cotab61